

Advances in the initialization of probit model estimation and the `{ino}` R package

Lennart Oelschläger Dietmar Bauer Marius Ötting

Bielefeld University, Econometrics Group

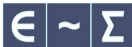
18 November 2022

What is this talk about?



- 1 The initialization problem
- 2 The probit model
- 3 New initialization idea
- 4 The `{ino}` R package

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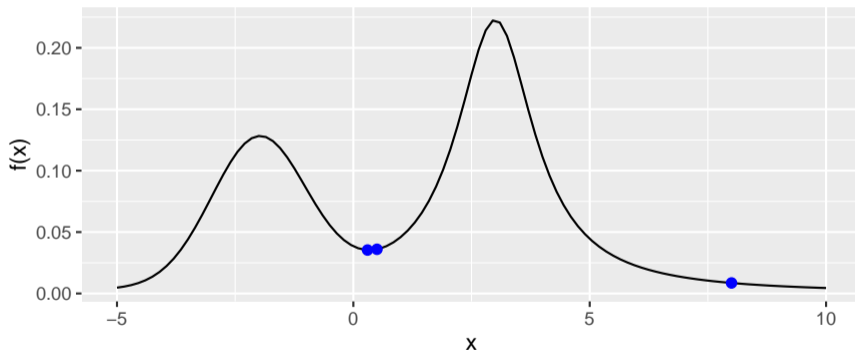
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Why do we have a problem?



Numerical optimization path

3 different starting points



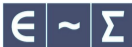
We find a local optimum



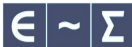
We find the global optimum



Convergence takes long



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Given choice data

- Discrete choice of decider n : $y_n \in \{1, \dots, J\}$
- Matrix of (alternative- or decider-specific) covariates of n : $X_n \in \mathbb{R}^{J \times P}$

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Probit model

$$U_n = X_n \beta + \epsilon_n \in \mathbb{R}^J \quad (\text{Latent utilities})$$

$$\epsilon_n \sim \mathcal{N}_J(0, \Sigma) \quad (\text{Error term})$$

$$y_n = \arg \max U_n \quad (\text{Choice link})$$

What we want? Estimates $\hat{\beta}$ (mean sensitivities) and $\hat{\Sigma}$ (error characterization).

The probit model (like any utility model) must be normalized:

Scale normalization

- $U > U' \Leftrightarrow c \cdot U > c \cdot U' \quad \forall c \in \mathbb{R}_+$
- For identification, fix, e.g., one entry of β to 1 (determines c)

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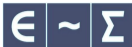
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Level normalization

- $U > U' \Leftrightarrow U + k > U' + k \quad \forall k \in \mathbb{R}$
- Consider utility differences: $U > U' \Leftrightarrow (U + k) - (U' + k) > 0$ (cancels k)
- Note: we loose one dimension ($J \rightsquigarrow J - 1$)

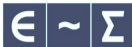
Utility differences



Difference utility vector $U_n \in \mathbb{R}^J$ with respect to some reference alternative i :

$$\Delta_i U_n \in \mathbb{R}^{J-1}$$

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$$\Delta_{y_n} U_n < 0.$$

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The difference operator looks like this:

$$\Delta_i = \begin{matrix} & & & & i \\ & & & & \\ & & & & \\ & & & & \\ i-1 & & & & \\ i+1 & & & & \end{matrix} \begin{pmatrix} 1 & & & & -1 \\ & \ddots & & & -1 & & & & 0 \\ & & & 1 & -1 & & & & \\ & & & & -1 & & 1 & & \\ 0 & & & & -1 & & & \ddots & \\ & & & & -1 & & & & 1 \end{pmatrix} \in \{-1, 0, 1\}^{(J-1) \times J}$$

Probability for choosing alternative i :

$$P_{ni}(\beta, \Sigma) = \text{Prob}(\Delta_i U_n < 0) = \underbrace{\Phi_{J-1}(-\Delta_i X_n \beta \mid 0, \Delta_i \Sigma \Delta_i')}_{\text{Computation expensive}}$$

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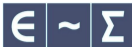
$$\log \mathcal{L}(\beta, \Sigma) = \sum_n \log P_{ny_n}(\beta, \Sigma)$$

Find MLE:

$$(\hat{\beta}, \hat{\Sigma}) = \arg \max \log \mathcal{L}(\beta, \Sigma)$$

Note: instead of Σ , optimize over L with $\Sigma = LL'$, where L is the lower-triangular Cholesky root with positive diagonal entries (for uniqueness)

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Using regression



Let's initialize β .

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1. Assume that Σ is known (if unknown, set $\Sigma = 1^{J \times J}$)
2. Consider first-order Taylor approximation of $P_{n\cdot}$ around 0:

$$P_{n\cdot}(-\Delta; X_n \beta \mid \Sigma) = P_{n\cdot}(0 \mid \Sigma) + \nabla P_{n\cdot}(0 \mid \Sigma) \cdot (-\Delta; X_n \beta) + R$$

3. Since $\mathbb{E}(y_n \mid \Sigma) = P_{n\cdot}(-\Delta; X_n \beta \mid \Sigma)$:

$$y_n = P_{n\cdot}(0 \mid \Sigma) + \underbrace{\nabla P_{n\cdot}(0 \mid \Sigma)}_{\tilde{X}_n} \cdot (-\Delta; X_n \beta) + e_n \quad (\text{not a catch-22!})$$

4. Compute OLS estimator $\hat{\beta}_{OLS}$ (very fast, just matrix product and inverting)

And what about Σ ?

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Trigger warning

Bayes people, please cover your eyes. Abuse of Bayes idea incoming.

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Idea:

1. Assume that β is known (if unknown, set $\beta = \hat{\beta}_{OLS}$)
2. Consider posterior of model parameters, including augmented $(U_n)_n$:

$$\text{Prob}(\beta, \Sigma, U | y) \propto \text{Prob}(\beta, \Sigma) \cdot \text{Prob}(U | \beta, \Sigma) \cdot 1\{y_n = \arg \max U_n\}$$

3. Assume conjugate prior and draw from posterior using Gibbs sampling (fairly fast)
4. Find $\hat{\Sigma}_{MCMC}$ as marginal posterior mode

Algorithm:

1. Initialize $\Sigma = 1^{J \times J}$
2. Estimate $\hat{\beta}_{OLS}$ using OLS
3. Estimate $\hat{\Sigma}_{MCMC}$ via MCMC
4. Initialize MLE with $(\hat{\beta}_{OLS}, \hat{\Sigma}_{MCMC})$

Hope:

- with Step 2 and 3, we initialize MLE close at the global optimum
- so that 4 is faster and more likely converges

Settings: $N = 200$, $J = 4$, $P = 4$, $X_n \stackrel{iid}{\sim} \mathcal{N}(0, 1)^{J \times P}$

True parameter: $\beta \sim \mathcal{N}(0, 1)^P$, $\beta_1 = 1$, $\Sigma = LL' \sim \mathcal{W}^{-1}$

Compare: Random initialization versus strategy in terms of computation time (sec) and deviation of MLE from true parameter (ndev)

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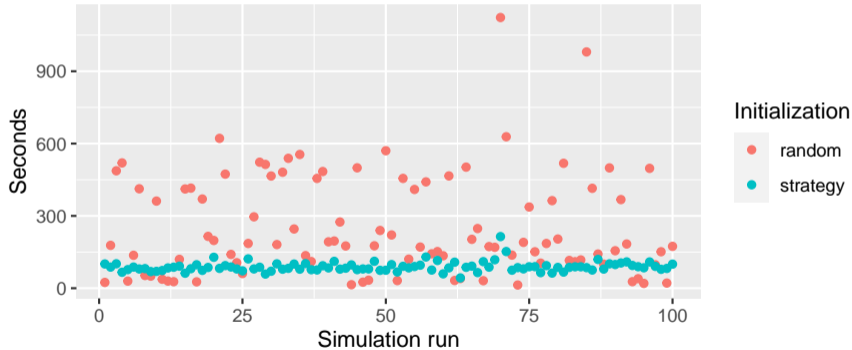
Compare: Random initialization versus strategy in terms of computation time (sec) and deviation of MLE from true parameter (ndev)

Table 1: One example run.

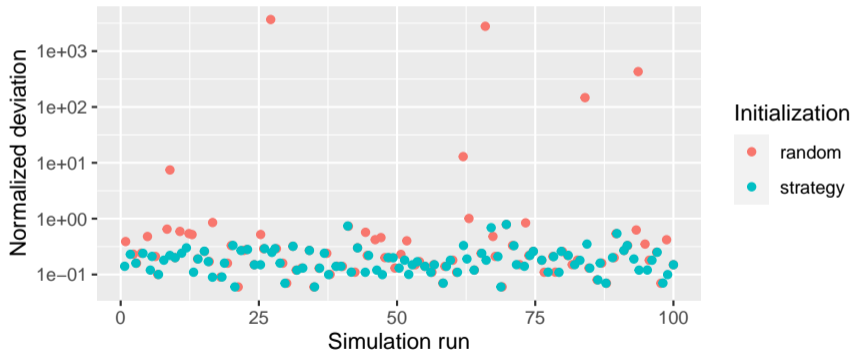
	b_1	b_2	b_3	b_4	l_1	l_2	l_3	l_4	l_5	l_6	sec	ndev
true_par	1	0.57	-1.10	1.17	2.46	-0.04	-0.13	2.38	0.10	1.13	0.00	0.00
init_random	1	-0.10	1.15	-0.91	0.23	-0.17	-0.91	0.78	-0.27	0.95	0.00	0.43
est_random	1	0.69	-1.14	1.36	2.10	-1.25	-0.32	1.48	0.38	0.92	487.14	0.16
init_strategy	1	0.87	-1.28	1.25	1.99	-1.22	-0.36	0.72	-0.52	0.66	1.20	0.23
est_strategy	1	0.69	-1.14	1.36	2.10	-1.25	-0.32	1.48	0.38	0.92	101.37	0.16

Improvement of computation time

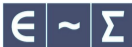
Strategy yields faster MLE in 79 of 100 cases



Improvement of convergence to global optimum
Strategy has same or smaller deviation in 98 of 100 cases



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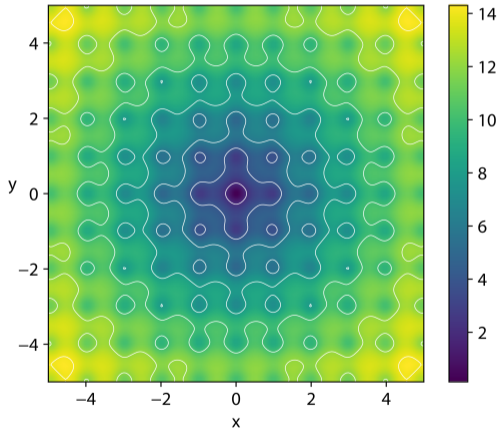
- Joint work with Marius
- Implements strategies for the initialization of numerical optimization:
 - effect of random initialization versus fixed initialization
 - effect of standardizing covariates
 - effect of subsetting covariates
 - effect of alternating optimization
 - comparing optimizer
 - number of identified optima
- Available on CRAN

```
> library("ino")
```



```
> x <- setup_ino(  
+   f = f_ackley,  
+   npar = 2,  
+   global = c(0,0),  
+   opt = set_optimizer_nlm()  
+ )
```

```
## Function to be optimized  
## f: f_ackley  
## npar: 2  
##  
## Numerical optimizer  
## 'stats::nlm': <optimizer 'stats::nlm'>  
##  
## Optimization runs  
## Records: 0
```



```
> random_initialization(x) %>% get_vars(vars = ".estimate")
```

```
## [1] 2.82139e-07 -1.75042e-07
```

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```
> x <- random_initialization(  
+ x, runs = 100, ncores = 3,  
+ sampler = function() stats::rnorm(npar(x))  
+ )
```

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## [1] 2.82139e-07 -1.75042e-07
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+ x, runs = 100, ncores = 3,  
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```
> overview_optima(x, digits = 2)
```

```
## optimum frequency  
## 1 0 44  
## 2 2.58 36  
## 3 3.57 12  
## 4 5.38 6  
## 5 6.56 1  
## 6 7.96 1
```

Thanks for listening!

Key message:

- MLE for probit model is sensitive to initial values
- Regression + MCMC reduce computation time
- $\{\text{ino}\}$ provides universal initialization strategies

Open questions:

- Consistency of strategy?
- How to initialize parameters of mixing distribution?

Please let me know:

- How is initialization an issue for you?
- Thoughts on $\{\text{ino}\}$?
- Other ideas for initialization?



loelschlaeger.de/talks



loelschlaeger@uni-bielefeld.de